Tutorial 7: General Random Variables 3

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 Two continuous random variables X, Y satisfying

$$P((X,Y) \in B) = \iint_{(x,y)\in B} f_{X,Y}(x,y)dxdy$$

for every subset B

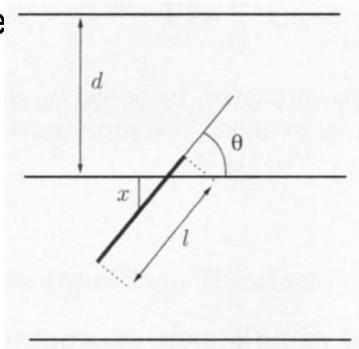
Revise - Marginal Probability

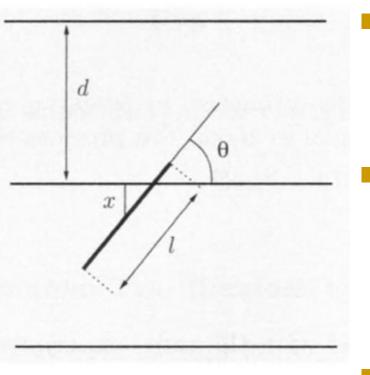
$$P(X \in A) = P(X \in A, Y \in (-\infty, \infty))$$
$$= \int_{A} \int_{-\infty}^{\infty} f_{X,Y}(x, y) dy dx$$
The mean is a DDE of *X* is

The marginal PDF of X is

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy$$

- A surface is ruled with parallel lines, which at distance d from each other.
- Suppose we throw a needle
- of length *l* randomly.
- What is the prob. that the needle will intersect one of the lines?

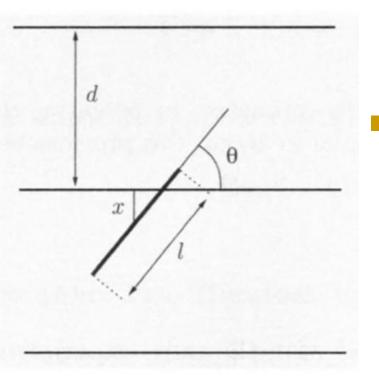




- Assume *l* < *d* so that the needle cannot intersect two lines simultaneously.
- X, the distance from the middle point of the needle and the nearest of the parallel lines
- ϑ , the acute angle formed by the needle and the lines

 We model (X, θ) with a uniform joint PDF so that

$$f_{X,\vartheta}(x,\theta) = \begin{cases} \frac{4}{\pi d}, & \text{if } x \in \left[0, \frac{d}{2}\right] \text{ and } \theta \in [0, \frac{\pi}{2}]\\ 0, & \text{otherwise} \end{cases}$$



The needle will intersect one of the lines if and only if $X \le \frac{l}{2} \sin \vartheta$

So the probability of intersection is

$$P\left(X \le \frac{l}{2}\sin\vartheta\right) = \iint_{X \le \frac{l}{2}\sin\theta} f_{X,\vartheta}(x,\theta) dx d\theta$$
$$= \frac{4}{\pi d} \int_0^{\pi/2} \int_0^{\left(\frac{l}{2}\right)\sin\theta} dx d\theta = \frac{4}{\pi d} \int_0^{\pi/2} \left(\frac{l}{2}\right)\sin\theta d\theta$$
$$= \frac{2l}{\pi d} (-\cos\theta) \left| \frac{\pi}{2} = \frac{2l}{\pi d} \right|_0^{\pi/2} = \frac{2l}{\pi d}$$

Let $X \in \mathbb{R}^n$ be a n-dimensional normal random variable. The PDF of X is then

$$f_X(x) = \frac{1}{(2\pi)^{\frac{n}{2}} |\Sigma|^{\frac{1}{2}}} e^{-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)}$$

The mean and variance

$$E(X) = \mu$$
$$V(X) = \Sigma$$

(Recall)

$$f_X(x) = \frac{1}{(2\pi)^{\frac{n}{2}} |\Sigma|^{\frac{1}{2}}} e^{-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)}$$

■ To draw: suppose we have Z_i~N(0,1) i.i.d and

$$Z = (Z_1, \dots, Z_n)$$

• We have $\mu + AZ \sim \mathbb{N}(\mu, \Sigma)$ where $AA^T = \Sigma$

Element-wise variance

$$\Sigma_{ij} = Cov(X_i, X_j)$$

Degenerate case

$$\Sigma = diag(v_1, v_2, \dots, v_n)$$

- Fact: in degenerate case, elements of X are independent
- Proof: As the coefficient of $x_i x_j$ is zero in $(x \mu)^T \Sigma^{-1} (x \mu)$, the PDF can be decomposed into the product of x_i part and x_j part.

 For conditional distribution: Gaussian variable condition on Gaussian variable is still Gaussian. We omit the proof.