# **ENGG2430A Probability and Statistics for Engineers**

# Chapter 2: Discrete Random Variables

#### Instructor: Shengyu Zhang

#### Content

#### Basic Concepts

- Probability Mass Function
- Functions of Random Variables
- Expectation, Mean, and Variance
- Joint PMFs of Multiple Random Variables
- Conditioning
- Independence

## Basic Concepts

In some experiments, the outcomes are numerical.

- □ E.g. stock price.
- In some other experiments, the outcomes are not numerical, but they may be associated with some numerical values of interest.
- *Example*. Selection of students from a given population, we may wish to consider their grade point average.
  - The students are not numerical, but their GPA scores are.

## Basic Concepts

- When dealing with these numerical values, it is useful to assign probabilities to them.
- This is done through the notion of a random variable.



# Main Concepts Related to Random Variables

- Starting with a probabilistic model of an experiment:
- A random variable is a real-valued function of the outcome of the experiment.
- A function of a random variable defines another random variable.



- 5 tosses of a coin.
- This is a random variable:

#### The number of heads

This is not:



# Main Concepts Related to Random Variables

- We can associate with each random variable certain "averages" of interest, such as the mean and the variance.
- A random variable can be conditioned on an event or on another random variable.
- Notion of independence of a random variable from an event or from another random variable.
- We'll talk about all these in this lecture.

#### Discrete Random Variable

- A random variable is called discrete if its range is either finite or countably infinite.
- *Example*. Two rolls of a die.
  - The sum of the two rolls.
  - The number of sixes in the two rolls.
  - □ The second roll raised to the fifth power.

#### Continuous random variable

- *Example*. Pick a real number a and associate to it the numerical value  $a^2$ .
- The random variable a<sup>2</sup> is continuous, not discrete.
- We'll talk about continuous random variables later.
- The following random variable is discrete:

$$sign(a) = \begin{cases} 1 & a > 0 \\ 0 & a = 0 \\ -1 & a < 0 \end{cases}$$

## Discrete Random Variables: Concepts

- A discrete random variable is a real-valued function of the outcome of a discrete experiment.
- A discrete random variable has an associated probability mass function (PMF), which gives the probability of each numerical value that the random variable can take.
- A function of a discrete random variable defines another discrete random variable, whose PMF can be obtained from the PMF of the original random variable.

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- For a discrete random variable X, the probability mass function (PMF) of X captures the probabilities of the values that it can take.
- If x is any possible value of X, the probability mass of x, denoted  $p_X(x)$ , is the probability of the event  $\{X = x\}$  consisting of all outcomes that give rise to a value of X equal to x :  $p_X(x) = P(\{X = x\})$



- Two independent tosses of a fair coin
- X: the number of heads obtained

# The PMF of X is $p_X(x) = \begin{cases} 1/4 & \text{if } x = 0 \text{ or } x = 2\\ 1/2 & \text{if } x = 1\\ 0 & \text{otherwise} \end{cases}$

Upper case characters to denote random variables

- $\square X, Y, Z, \dots$
- Lower case characters to denote real numbers
  - $\square x, y, z, \dots$
  - the numerical values of a random variable
- We'll write P(X = x) in place of the notation  $P({X = x})$ .
- Similarly, we'll write  $P(X \in S)$  for the probability that X takes a value within a set S.

 Follows from the additivity and normalization axioms

$$\sum_{\substack{x: \ all \ possible \\ values \ of \ X}} p_X(x) = 1$$

- The events  $\{X = x\}$  are disjoint, and they form a partition of the sample space
- For any set S of real numbers

$$P(X \in S) = \sum_{x \in S} p_X(x)$$

For each possible value x of X:

• Collect all the possible outcomes that give rise to the event  $\{X = x\}$ .

• Add their probabilities to obtain  $p_X(x)$ .



Important specific distributions

- Binomial random variable
- Geometric random variable
- Poisson random variable

#### Bernoulli Random Variable

The Bernoulli random variable takes the two values 1 and 0

 $X \in \{0,1\}$ 

Its PMF is

$$p_X(x) = \begin{cases} p & \text{if } x = 1\\ 1 - p & \text{if } x = 0 \end{cases}$$

Example of Bernoulli Random Variable

The state of a telephone at a given time that can be either free or busy.

A person who can be either healthy or sick with a certain disease.

The preference of a person who can be either for or against a certain political candidate.

#### The Binomial Random Variable

A biased coin is tossed n times.

- Each toss is independently of prior tosses
   Head with probability *p*.
   Tail with probability 1 *p*.
- The number X of heads up is a binomial random variable.

#### The Binomial Random Variable

We refer to X as a binomial random variable with parameters n and p.

For 
$$k = 0, 1, ..., n$$
.  
 $p_X(k) = P(X = k) = \binom{n}{k} p^k (1 - p)^{n-k}$ 

#### The Binomial Random Variable





#### The Geometric Random Variable

- Independently and repeatedly toss a biased coin with probability of a head p, where 0
- The geometric random variable is the number X of tosses needed for a head to come up for the first time.

## The Geometric Random Variable

The PMF of a geometric random variable

$$p_X(k) = (1-p)^{k-1}p$$

• k - 1 tails followed by a head.

Normalization condition is satisfied:

$$\sum_{k=1}^{\infty} p_X(k) = \sum_{k=1}^{\infty} (1-p)^{k-1} p = p \sum_{k=0}^{\infty} (1-p)^k$$
$$= p \cdot \frac{1}{1-(1-p)} = 1$$

The Geometric Random Variable

• The  $p_X(k) = (1-p)^{k-1}p$  decreases as a geometric progression with parameter 1-p.



#### The Poisson Random Variable

 A Poisson random variable takes nonnegative integer values.
 The PMF

$$p_X(k) = e^{-\lambda} \frac{\lambda^k}{k!}, \quad k = 0, 1, 2, ...,$$

Normalization condition

$$\sum_{k=0}^{\infty} e^{-\lambda} \frac{\lambda^k}{k!} = e^{-\lambda} \left( 1 + \lambda + \frac{\lambda^2}{2!} + \frac{\lambda^3}{3!} + \cdots \right)$$
$$= e^{-\lambda} e^{\lambda} = 1$$

- Poisson random variable can be viewed as a binomial random variable with very small p and very large n.
- More precisely, the Poisson PMF with parameter  $\lambda$  is a good approximation for a binomial PMF with parameters n and p where  $\lambda = np$ , n is large and p is small.
  - □ See the <u>wiki page</u> for a proof.

# Examples

- Because of the above connection, Poisson random variables are used in many scenarios.
- X is the number of typos in a book of n words.
   The probability that any one word is misspelled is very
  - small.
- X is the number of cars involved in accidents in a city on a given day.
  - The probability that any one car is involved in an accident is very small.

#### The Poisson Random Variable

- For Poisson random variable  $p_X(k) = e^{-\lambda} \frac{\lambda^{\kappa}}{k!}$ 
  - λ ≤ 1, monotonically decreasing
     λ > 1, first increases and then decreases



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#### Functions of Random Variables

- Consider a probability model of today's weather
  - X = the temperature in degrees Celsius
  - Y = the temperature in degrees Fahrenheit
- Their relation is given by

Y = 1.8X + 32

In this example, Y is a linear function of X, of the form

$$Y = g(X) = aX + b$$

#### Functions of Random Variables

We may also consider nonlinear functions, such as

$$Y = \log(X)$$

- In general, if Y = g(X) is a function of a random variable X, then Y is also a random variable.
- The PMF  $p_Y$  of Y = g(X) can be calculated from PMF  $p_X$  of X

$$p_Y(y) = \sum_{x:g(x)=y} p_X(x)$$



#### The PMF of X is

$$p_X(x) = \begin{cases} 1/9 & \text{if } x \text{ is an integer and } x \in [-4,4] \\ 0 & \text{otherwise} \end{cases}$$

• Let 
$$Y = |X|$$
. Then the PMF of Y is

$$p_Y(y) = \begin{cases} 2/9 & \text{if } y = 1,2,3,4 \\ 1/9 & \text{if } y = 0 \\ 0 & \text{otherwise} \end{cases}$$



#### Visualization of the relation between X and Y





#### • Let $Z = X^2$ . Then the PMF of Z is

$$p_Z(z) = \begin{cases} 2/9 & \text{if } z = 1, 4, 9, 16\\ 1/9 & \text{if } z = 0\\ 0 & \text{otherwise} \end{cases}$$

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#### Expectation

- Sometimes it is desirable to summarize the values and probabilities by one number.
- The expectation of X is a weighted average of the possible values of X.
  - Weights: probabilities.
- Formally, the expected value of a random variable X, with PMF  $p_X(x)$ , is

 $\mathbf{E}[X] = \sum_{x} x p_X(x)$ 

Names: expected value, expectation, mean

## Example

Two independent coin tosses

• 
$$P(H) = \frac{3}{4}$$
  
•  $X =$  the number of heads

Binomial random variable with parameters n = 2 and p = 3/4.



#### The PMF is

$$p_X(k) = \begin{cases} (1/4)^2 & \text{if } k = 0\\ 2 \cdot (1/4) \cdot (3/4) & \text{if } k = 1\\ (3/4)^2 & \text{if } k = 2 \end{cases}$$

The mean is  

$$\mathbf{E}[X] = 0 \cdot \left(\frac{1}{4}\right)^2 + 1 \cdot \left(2 \cdot \frac{1}{4} \cdot \frac{3}{4}\right) + 2 \cdot \left(\frac{3}{4}\right)^2 = \frac{3}{2}$$

#### Expectation

Consider the mean as the center of gravity of the PMF



#### Variance

- Besides the mean, there are several other important quantities.
- The *k*th moment is  $\mathbf{E}[X^k]$

So the first moment is just the mean.

- Variance of X, denoted by var(X), is  $var(X) = \mathbf{E}[(X - \mathbf{E}[X])^2]$ 
  - The second moment of  $X \mathbf{E}[X]$ .
- The variance is always non-negative:  $var(X) \ge 0$

#### Standard deviation

- Variance is closely related to another measure.
- Standard deviation of *X*, denoted by  $\sigma_X$ , is  $\sigma_X = \sqrt{\operatorname{var}(X)}$



Suppose that the PMF of X is

 $p_X(x) = \begin{cases} 1/9 & \text{if } x \text{ is an integer and } x \in [-4,4] \\ 0 & \text{otherwise} \end{cases}$ 

The expectation

$$\mathbf{E}[X] = \sum_{x} x p_X(x) = \frac{1}{9} \sum_{x=-4}^{4} x = 0$$

Can also be seen from symmetry.



• Let 
$$Z = (X - \mathbf{E}[X])^2 = X^2$$
. The PMF of Z  
 $p_Z(z) = \begin{cases} 2/9 & \text{if } z = 1, 4, 9, 16 \\ 1/9 & \text{if } z = 0 \\ 0 & \text{otherwise} \end{cases}$ 

• The variance of X is then  $var(X) = \mathbf{E}[Z] = \sum_{Z} zp_{Z}(Z)$  $= 0 \cdot \frac{1}{9} + 1 \cdot \frac{2}{9} + 4 \cdot \frac{2}{9} + 9 \cdot \frac{2}{9} + 16 \cdot \frac{2}{9} = \frac{60}{9}$ 

## Expectation for g(X)

- There is a simpler way of computing var(g(X)).
- Let X be a random variable with PMF  $p_X(x)$ , and let g(X) be a real-valued function of X.
- The expected value of the random variable Y = g(X) is

$$\mathbf{E}[g(X)] = \sum_{x} g(x) p_X(x)$$

#### Expectation for g(X)

• Using the formula  $p_Y(y) = \sum_{\{x \mid g(x) = y\}} p_X(x)$ :  $\mathbf{E}[q(X)] = \mathbf{E}[Y]$  $=\sum_{v} y p_{Y}(y)$  $= \sum_{y} y \sum_{\{x \mid q(x) = y\}} p_X(x)$  $= \sum_{y} \sum_{\{x \mid g(x)=y\}} y p_X(x)$  $= \sum_{y} \sum_{\{x \mid g(x)=y\}} g(x) p_X(x)$  $=\sum_{x} q(x) p_{x}(x)$ 

## Variance example

#### The PMF of X

 $p_X(x) = \begin{cases} 1/9 & \text{if } x \text{ is an integer and } x \in [-4,4] \\ 0 & \text{otherwise} \end{cases}$ 

The variance

$$var(X) = \mathbf{E}[(X - \mathbf{E}[X])^{2}]$$
  
=  $\sum_{x} (x - \mathbf{E}[X])^{2} p_{X}(x)$   
=  $\frac{1}{9} \sum_{x=-4}^{4} x^{2}$   
=  $(16 + 9 + 4 + 1 + 0 + 1 + 9 + 16)/9$   
=  $\frac{60}{9}$ 

#### Mean of aX + b

#### • Let *Y* be a linear function of *X* Y = aX + b

The mean of Y

$$\mathbf{E}[Y] = \sum_{x} (ax + b)p_X(x)$$
$$= a \sum_{x} x p_X(x) + b \sum_{x} p_X(x) = a\mathbf{E}[X] + b$$

The expectation scales *linearly*.

#### Variance of aX + b

#### • Let *Y* be a linear function of *X* Y = aX + b

#### The variance of Y

$$\operatorname{var}(Y) = \sum_{x} (ax + b - \mathbf{E}[aX + b])^{2} p_{X}(x)$$
$$= \sum_{x} (ax + b - a\mathbf{E}[X] - b)^{2} p_{X}(x)$$
$$= a^{2} \sum_{x} (x - \mathbf{E}[X])^{2} p_{X}(x)$$
$$= a^{2} \operatorname{var}(X)$$

The variance scales *quadratically*.

#### Variance as moments

• Fact. 
$$var(X) = \mathbf{E}[X^2] - (\mathbf{E}[X])^2$$
.

• 
$$var(X) = \mathbf{E}[(X - \mathbf{E}[X])^2]$$
  
 $= \mathbf{E}[X^2 - 2X\mathbf{E}[X] + (\mathbf{E}[X])^2]$   
 $= \mathbf{E}[X^2] - 2\mathbf{E}[X\mathbf{E}[X]] + (\mathbf{E}[X])^2$   
 $= \mathbf{E}[X^2] - 2\mathbf{E}[X]\mathbf{E}[X] + (\mathbf{E}[X])^2$   
 $= \mathbf{E}[X^2] - (\mathbf{E}[X])^2$ 

#### Example: Average time

- Distance between class and home is 2 miles
- P(weather is good) = 0.6
- Speed:
  - $\Box$  V = 5 miles/hour if weather is good.
  - $\Box$  V = 30 miles/hour if weather is bad.
- Question: What is the mean of the time T to get to class?

Example: Average time

## The PMF of T $p_T(t) = \begin{cases} 0.6 & \text{if } t = \frac{2}{5} \text{ hours} \\ 0.4 & \text{if } t = \frac{2}{30} \text{ hours} \end{cases}$

• The mean of T  

$$\mathbf{E}[T] = 0.6 \cdot \frac{2}{5} + 0.4 \cdot \frac{2}{30} = \frac{4}{15}$$

## Example: Average time

Wrong calculation by speed V The mean of speed V  $\mathbf{E}[V] = 0.6 \cdot 5 + 0.4 \cdot 30 = 15$ The mean of time T  $\overline{\mathbf{E}[V]} = \overline{15}$ To summarize, in this example we have  $T = \frac{2}{v}$  and  $\mathbf{E}[T] = \mathbf{E} \left| \frac{2}{v} \right| \neq \frac{2}{\mathbf{E}[v]}$ 

#### Example: Bernoulli

Consider the Bernoulli random variable X with PMF

$$p_X(x) = \begin{cases} p & \text{if } x = 1\\ 1 - p & \text{if } x = 0 \end{cases}$$

• Its mean, second moment, and variance:  $E[X] = 1 \cdot p + 0 \cdot (1 - p) = p$   $E[X^2] = 1^2 \cdot p + 0 \cdot (1 - p) = p$  $var(X) = E[X^2] - (E[X])^2 = p - p^2 = p(1 - p)$ 

#### Example: Uniform

What is the mean and variance of the roll of a fair six-sided die?

$$p_X(k) = \begin{cases} 1/6 & \text{if } k = 1,2,3,4,5,6 \\ 0 & \text{otherwise} \end{cases}$$

• The mean  $\mathbf{E}[X] = 3.5$  and the variance  $\operatorname{var}(X) = \mathbf{E}[X^2] - (\mathbf{E}[X])^2$   $= \frac{1}{6} (1^2 + 2^2 + 3^2 + 4^2 + 5^2 + 6^2) - 3.5^2$ = 35/12

- General, a discrete uniformly distributed random variable
  - Range: contiguous integer values a, a + 1, ..., b
     Probability: equal probability
- The PMF is

$$p_X(k) = \begin{cases} \frac{1}{b-a+1} & \text{if } k = a, a+1, \dots, b\\ 0 & \text{otherwise} \end{cases}$$

The mean

$$\mathbf{E}[X] = \frac{a+b}{2}$$

- For variance, first consider a = 1 and b = n
- The second moment

$$\mathbf{E}[X^2] = \frac{1}{n} \sum_{k=1}^n k^2 = \frac{1}{6} (n+1)(2n+1)$$

# • The variance for special case $var(X) = \mathbf{E}[X^2] - (\mathbf{E}[X])^2$ $= \frac{1}{6}(n+1)(2n+1) - \frac{1}{4}(n+1)^2$ $= \frac{n^2 - 1}{12}$

- For the case of general integers a and b
  - X: discrete uniform over [a, b]
  - *Y*: discrete uniform over [1, b a + 1]
- Relation between X and Y

$$Y = X - a + 1$$

Thus

$$var(X) = var(Y) = \frac{(b-a+1)^2 - 1}{12}$$

## Example: Poisson

#### Recall Poisson PMF

$$p_X(k) = e^{-\lambda} \frac{\lambda^k}{k!}$$
  $k = 0, 1, 2, ...,$ 

Mean:

$$\mathbf{E}[X] = \sum_{k=0}^{\infty} k e^{-\lambda} \frac{\lambda^k}{k!} = \sum_{k=1}^{\infty} k e^{-\lambda} \frac{\lambda^k}{k!}$$
$$= \lambda \sum_{k=1}^{\infty} e^{-\lambda} \frac{\lambda^{k-1}}{(k-1)!} = \lambda \sum_{m=0}^{\infty} e^{-\lambda} \frac{\lambda^m}{m!}$$
$$= \lambda$$

- Variance:  $var[X] = \lambda$ .
  - Verification left as exercise.

- A person is given two questions and must decide which question to answer first.
  - P(question 1 correct) = 0.8 Prize=\$100
  - P(question 2 correct) = 0.5 Prize=\$200
  - If incorrectly answer the first question, then no second question.
- How to choose the first question so that maximize the expected prize?

#### Tree illustration



• Answer question 1 first: Then the PMF of X is  $p_X(0) = 0.2$   $p_X(100) = 0.8 \cdot 0.5$  $p_X(300) = 0.8 \cdot 0.5$ 

• We have  $\mathbf{E}[X] = 0.8 \cdot 0.5 \cdot 100 + 0.8 \cdot 0.5 \cdot 300 = 160$ 

Answer question 2 first: Then the PMF of X is

$$p_X(0) = 0.5$$
  
 $p_X(200) = 0.5 \cdot 0.2$   
 $p_X(300) = 0.5 \cdot 0.8$ 

- We have  $\mathbf{E}[X] = 0.5 \cdot 0.2 \cdot 200 + 0.5 \cdot 0.8 \cdot 300 = 140$
- It is better to answer question 1 first.

- Let us now generalize the analysis.
  - $p_1$ : P(correctly answering question 1)
  - □  $p_2$ : *P*(correctly answering question 2)
  - $v_1$ : prize for question 1
  - $v_2$ : prize for question 2

• Answer question 1 first  $E[X] = p_1(1 - p_2)v_1 + p_1p_2(v_1 + v_2)$   $= p_1v_1 + p_1p_2v_2$ 

## • Answer question 2 first $E[X] = p_2(1 - p_1)v_2 + p_2p_1(v_2 + v_1)$ $= p_2v_2 + p_2p_1v_1$

It is optimal to answer question 1 first if and only if

 $p_1v_1 + p_1p_2v_2 \ge p_2v_2 + p_2p_1v_1$ 

Or equivalently

$$\frac{p_1 v_1}{1 - p_1} \ge \frac{p_2 v_2}{1 - p_2}$$

Rule: Order the questions in decreasing value of the expression  $\frac{pv}{(1-p)}$ 

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#### Multiple Random Variables

Probabilistic models often involve several random variables of interest.

- Example: In a medical diagnosis context, the results of several tests may be significant.
- Example: In a networking context, the workloads of several gateways may be of interest.

#### Joint PMFs of Multiple Random Variables

- Consider two discrete random variables X and Y associated with the same experiment.
- The joint PMF of X and Y is denoted by p<sub>X,Y</sub>. It specifies the probability of the values that X and Y can take.
- If (x, y) is a pair of values that (X, Y) can take, then the probability mass of (x, y) is the probability of the event  $\{X = x, Y = y\}$ :  $P_{X,Y}(x, y) = P(X = x, Y = y).$

The joint PMF determines the probability of any event that can be specified in terms of the random variables X and Y.

• For example, if *A* is the set of all pairs (x, y) that have a certain property, then  $P((X, Y) \in A) = \sum_{(x, y) \in A} p_{X,Y}(x, y)$ 

Joint PMFs of Multiple Random Variables

The PMFs of X and Y  $p_X(x) = \sum_y p_{X,Y}(x, y), \quad p_Y(y) = \sum_x p_{X,Y}(x, y)$ The formula can be verified by  $p_X(x) = P(X = x)$   $= \sum_y P(X = x, Y = y)$   $= \sum_y p_{X,Y}(x, y)$ 

•  $p_X$ ,  $p_Y$  are the marginal PMFs.
### Joint PMFs of Multiple Random Variables

- Computing the marginal MPFs  $p_X$  and  $p_Y$  of  $p_{X,Y}$  from table.
- The joint PMF p<sub>X,Y</sub> is arranged in a twodimensional table.



### Joint PMFs of Multiple Random Variables

The marginal PMF of *X* or *Y* at a given value is obtained by adding the table entries along a corresponding column or row, respectively.



#### Functions of Multiple Random Variables

 One can generate new random variables by applying functions on several random variables.

• Consider 
$$Z = g(X, Y)$$
.

Its PMF can be calculated from the joint PMF p<sub>X,Y</sub> according to

$$p_Z(z) = \sum_{\{(x,y)|g(x,y)=z\}} p_{X,Y}(x,y)$$

### Functions of Multiple Random Variables

The expected value rule for multiple variables

$$\mathbf{E}[g(X,Y)] = \sum_{x,y} g(x,y) p_{X,Y}(x,y)$$

For special case, g is linear and of the form aX + bY + c, we have E[aX + bY + c] = aE[X] + bE[Y] + c
"linearity of expectation" --- regardless of dependence of X and Y.

## More than Two Random Variables

- We can also consider three or more random variables.
- The joint PMF of three random variables X, Y, and Z

$$p_{X,Y,Z}(x, y, z) = P(X = x, Y = y, Z = z)$$

The marginal PMFs are

$$p_{X,Y}(x,y) = \sum_{z} p_{X,Y,Z}(x,y,z)$$

and

$$p_X(x) = \sum_{y} \sum_{z} p_{X,Y,Z}(x, y, z)$$

### More than Two Random Variables

The expected value rule for functions
E[g(X,Y,Z)] = ∑<sub>x,y,z</sub> g(x,y,z)p<sub>X,Y,Z</sub>(x,y,z)
If g is linear and of the form
g(X,Y,Z) = aX + bY + cZ + d
then

$$\mathbf{E}[aX + bY + cZ + d]$$
  
=  $a\mathbf{E}[X] + b\mathbf{E}[Y] + c\mathbf{E}[Z] + d$ 

### More than Two Random Variables

Generalization to more than three random variables.

For any random variables  $X_1, X_2, \dots, X_n$  and any scalars  $a_1, a_2, \dots, a_n$ , we have  $\mathbf{E}[a_1X_1 + a_2X_2 + \dots + a_nX_n]$  $= a_1\mathbf{E}[X_1] + a_2\mathbf{E}[X_2] + \dots + a_n\mathbf{E}[X_n]$ 

- 300 students in probability class
- Each student has probability 1/3 of getting an A, independently of any other student.
- X: the number of students that get an A.
- Question: What is the mean of X?

- Let  $X_i$  be the random variable for *i*th student  $X_i = \begin{cases} 1 & \text{if the } i\text{th student gets an A} \\ 0 & \text{otherwise} \end{cases}$
- Each X<sub>i</sub> is a Bernoulli random variable
   E[X<sub>i</sub>] = p = 1/3
   Var[X<sub>i</sub>] = p(1-p) = (1/3)(2/3) = 2/9

The random variable X can be expressed as their sum

$$X = X_1 + X_2 + \dots + X_n$$

• Using the linearity of *X* as a function of the *X<sub>i</sub>*  $\mathbf{E}[X] = \sum_{i=1}^{300} \mathbf{E}[X_i] = \sum_{i=1}^{300} \frac{1}{3} = 300 \cdot \frac{1}{3} = 100$ 

If we repeat this calculation for a general number of students n and probability of A equal to p, we obtain

$$E[X] = \sum_{i=1}^{n} E[X_i] = np$$

## Example: The Hat Problem

- Suppose that n people throw their hats in a box.
- Each picks up one hat at random.
- X: the number of people that get back their own hat

Question: What is the expected value of X?

## Example: The Hat Problem

For the *i*th person, we introduce a random variable X<sub>i</sub>

$$X_i = \begin{cases} 1 & \text{if the } i\text{th his own} \\ 0 & \text{otherwise} \end{cases}$$

• Since 
$$P(X_i = 1) = \frac{1}{n}$$
 and  $P(X_i = 0) = 1 - \frac{1}{n}$   
 $E[X_i] = 1 \cdot \frac{1}{n} + 0 \cdot \left(1 - \frac{1}{n}\right) = \frac{1}{n}$ 

## Example: The Hat Problem

We know

$$X = X_1 + X_2 + \dots + X_n$$

## • Thus $\mathbf{E}[X] = \mathbf{E}[X_1] + \mathbf{E}[X_2] + \dots + \mathbf{E}[X_n] = n \cdot \frac{1}{n} = 1$

Summary of Facts About Joint PMFs

The joint PMF of X and Y is defined by

 $p_{X,Y}(x,y) = P(X = x, Y = y)$ 

The marginal PMFs of X and Y can be obtained from the joint PMF, using the formulas

 $p_X(x) = \sum_y p_{X,Y}(x,y), \ p_Y(y) = \sum_x p_{X,Y}(x,y)$ 

Summary of Facts About Joint PMFs

A function g(X,Y) of X and Y defines another random variable

$$\mathbf{E}[g(X,Y)] = \sum_{x,y} g(x,y) p_{X,Y}(x,y)$$

If g is linear, of the form aX + bY + c,  $\mathbf{E}[aX + bY + c] = a\mathbf{E}[X] + b\mathbf{E}[Y] + c$ 

These naturally extend to more than two random variables.

## Content

- Basic Concepts
- Probability Mass Function
- Functions of Random Variables
- Expectation, Mean, and Variance
- Joint PMFs of Multiple Random Variables
- Conditioning
- Independence

## Conditioning

- In a probabilistic model, a certain event A has occurred
- Conditional probability captures this knowledge.
- Conditional probabilities are like ordinary probabilities (satisfy the three axioms) except
  - refer to a new universe: event A is known to have occurred

The conditional PMF of a random variable X, conditioned on a particular event A with P(A) > 0, is defined by

$$p_{X|A}(x) = P(X = x|A)$$
$$= \frac{P(\{X = x\} \cap A)}{P(A)}$$

- Consider the events  $\{X = x\} \cap A$ :
  - They are disjoint for different values of x.
  - Their union is A.
- Thus  $P(A) = \sum_{x} P(\{X = x\} \cap A)$
- Combining this and  $p_{X|A}(x) = P(\{X = x\} \cap A)/P(A)$  (last slide), we can see that

 $\sum_{x} p_{X|A}(x) = 1$ 

• So  $p_{X|A}$  is a legitimate PMF.

- The conditional PMF is calculated similar to its unconditional counterpart.
- To obtain  $p_{X|A}(x)$ 
  - Add the probabilities of the outcomes X = x
  - Conditioning event A
  - Normalize by dividing with P(A)

Visualization and calculation of the conditional PMF  $p_{X|A}(x)$ 



## Example: dice

## X: the roll of a fair 6-sided dice A: the roll is an even number $p_{X|A}(x) = P(X = x|A)$ $=\frac{P(X=x \text{ and } A)}{P(A)}$ $= \begin{cases} \frac{1}{3} & \text{if } x = 2,4,6 \\ 0 & \text{otherwise} \end{cases}$

- We have talked about conditioning a random variable X on an event A.
- Now let's consider conditioning a random variable X on another random variable Y.
- Let X and Y be two random variables associated with the same experiment.
- The experimental value Y = y ( $p_Y(y) > 0$ ) provides partial knowledge about the value of *X*.

- The knowledge is captured by the conditional PMF  $p_{X|Y}$  of X given Y, which is defined as  $p_{X|A}$  for  $A = \{Y = y\}$ :  $p_{X|Y}(x|y) = P(X = x|Y = y)$
- Using the definition of conditional probabilities

$$p_{X|Y}(x|y) = \frac{P(X = x, Y = y)}{P(Y = y)} = \frac{p_{X,Y}(x, y)}{p_Y(y)}$$

- Fix some y, with  $p_Y(y) > 0$  and consider  $p_{X|Y}(x|y)$  as a function of x.
- This function is a valid PMF for X:
  - Assigns nonnegative values to each possible x
  - These values add to 1
  - Has the same shape as  $p_{X,Y}(x, y)$
  - $\Box \sum_{x} p_{X|Y}(x|y) = 1$

Visualization of the conditional PMF  $p_{X|Y}(x|y)$ 



It is convenient to calculate the joint PMF by a sequential approach and the formula  $p_{X,Y}(x,y) = p_Y(y)p_{X|Y}(x|y),$ 

Or its counterpart

$$p_{X,Y}(x,y) = p_X(x)p_{Y|X}(y|x).$$

This method is entirely similar to the use of the multiplication rule from previous lectures.

- A professor independently answers each of her students' questions incorrectly with probability 1/4.
- In each lecture the professor is asked 0,1, or 2 questions with equal probability 1/3.
  - *X*: the number of questions professor is asked
  - Y: the number of questions she answers wrong in a given lecture

- Construct the joint PMF  $p_{X,Y}(x, y)$ : calcualte all the probabilities P(X = x, Y = y).
- Using a sequential description of the experiment and the multiplication rule  $p_{X,Y}(x,y) = p_Y(y)p_{X|Y}(x|y)$



#### For example,

 $p_{X,Y}(1,1) = p_X(x)p_{Y|X}(y,x) = \frac{1}{4} \cdot \frac{1}{3} = \frac{1}{12}$ 

- We can compute other useful information from two-dimensional table.
- For example,

*P*(at least one wrong answer)

$$= p_{X,Y}(1,1) + p_{X,Y}(2,1) + p_{X,Y}(2,2)$$

$$=\frac{4}{48} + \frac{6}{48} + \frac{1}{48} = \frac{11}{48}$$

The conditional PMF can also be used to calculate the marginal PMFs.

$$p_X(x) = \sum_y p_{X,Y}(x,y) = \sum_y p_Y(y) p_{X|Y}(x|y)$$

This formula provides a divide-and-conquer method for calculating marginal PMFs.

# Summary of Facts About Conditional PMFs

- Conditional PMFs are similar to ordinary PMFs, but refer to a universe where the conditioning event is known to have occurred.
- The conditional PMF of X given an event A with P(A) > 0, is defined by  $p_{X|A}(x) = P(X = x|A)$

and satisfies

$$\sum_{x} p_{X|A}(x) = 1$$

## Summary of Facts About Conditional PMFs

The conditional PMF of X given Y can be used to calculate the marginal PMFs with the formula

$$p_X(x) = \sum_{y} p_Y(y) p_{X|Y}(x|y)$$

This is analogous to the divide-and-conquer approach for calculating probabilities using the total probability theorem.

## Conditional Expectations

The conditional expectation of X given an event A with P(A) > 0, is defined by

$$\mathbf{E}[X|A] = \sum_{x} x p_{X|A}(x|A)$$
  
For a function  $g(X)$ , it is given by  
$$\mathbf{E}[g(X)|A] = \sum_{x} g(x) p_{X|A}(x|A)$$
# Conditional Expectations

The conditional expectation of X given a value y of Y is defined by

$$\mathbf{E}[X|Y=y] = \sum_{x} x p_{X|Y}(x|y)$$

The total expectation theorem

$$\boldsymbol{E}[X] = \sum_{y} p_{Y}(y) \, \mathbf{E}[X|Y = y]$$

#### Conditional Expectations

• Let  $A_1, ..., A_n$  be disjoint events that form a partition of the sample space, and assume that  $P(A_i) > 0$  for all *i*. Then

 $\mathbf{E}[X] = \sum_{i=1}^{n} P(A_i) \mathbf{E}[X|A_i]$ 

Indeed,

$$E[X] = \sum_{x} x p_{X}(x)$$
  
=  $\sum_{x} x \sum_{i=1}^{n} P(A_{i}) p_{x|A_{i}}(x|A_{i})$   
=  $\sum_{i=1}^{n} P(A_{i}) \sum_{x} x p_{x|A_{i}}(x|A_{i})$   
=  $\sum_{i=1}^{n} P(A_{i}) \mathbf{E}[X|A_{i}]$ 

### Conditional Expectation

- Messages transmitted by a computer in Boston through a data network are destined
  - □ for New York with probability 0.5
  - for Chicago with probability 0.3
  - □ for San Francisco with probability 0.2
- The transit time X of a message is random
  - $\mathbf{E}[X] = 0.05$  for New York
  - $\mathbf{E}[X] = 0.1$  for Chicago
  - $\mathbf{E}[X] = 0.3$  for San Francisco

#### Conditional Expectation

#### • By total expectation theorem $E[X] = 0.5 \cdot 0.05 + 0.3 \cdot 0.1 + 0.2 \cdot 0.3$ = 0.115

- You write a software program over and over,
  - $\square$  probability p that it works correctly
  - independently from previous attempts
- X: the number of tries until the program works correctly
- Question: What is the mean and variance of X?

 X is a geometric random variable with PMF p<sub>X</sub>(k) = (1 - p)<sup>k-1</sup>p k = 1,2,...
 The mean and variance of X

$$\mathbf{E}[X] = \sum_{k=1}^{\infty} k(1-p)^{k-1}p$$
  
var(X) =  $\sum_{k=1}^{\infty} (k - \mathbf{E}[X])^2 (1-p)^{k-1}p$ 

- Evaluating these infinite sums is somewhat tedious.
- As an alternative, we will apply the total expectation theorem.

#### Let

$$A_1 = \{X = 1\} = \{\text{first try is a success}\}$$
  
and

- If the first try is successful, we have X = 1 $\mathbf{E}[X|X = 1] = 1$
- If the first try fails (X > 1), we have wasted one try, and we are back where we started.
  The expected number of remaining tries is E[X]
  We have

$$\mathbf{E}[X|X > 1] = 1 + \mathbf{E}[X]$$

- Thus
  - $\mathbf{E}[X]$
  - $= P(X = 1)\mathbf{E}[X|X = 1] + P(X > 1)\mathbf{E}[X|X > 1]$ = p + (1 - p)(1 + **E**[X])
- Solving this equation gives  $\mathbf{E}[X] = \frac{1}{p}$

# Similar reasoning $\mathbf{E}[X^2 | X = 1] = 1$

and

$$\mathbf{E}[X^{2}|X > 1] = \mathbf{E}[(1+X)^{2}]$$
$$= 1 + 2\mathbf{E}[X] + \mathbf{E}[X^{2}]$$

• So  $\mathbf{E}[X^2] = p \cdot 1 + (1-p)(1+2\mathbf{E}[X] + \mathbf{E}[X^2])$ 

We obtain

$$\mathbf{E}[X^2] = \frac{2}{p^2} - \frac{1}{p}$$

• and conclude that  $Var(X) = E[X^{2}] - (E[X])^{2}$   $= \frac{2}{p^{2}} - \frac{1}{p} - \frac{1}{p^{2}} = \frac{1-p}{p^{2}}$ 

#### Content

- Basic Concepts
- Probability Mass Function
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- Joint PMFs of Multiple Random Variables
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- Independence

- Idea is similar to the independence of two events.
- Knowing the occurrence of the conditioning event tells us nothing about the value of the random variable.

• Formally, the random variable X is independent of the event A if  $P(X = x \text{ and } A) = P(X = x)P(A) = p_X(x)P(A)$ 

Same as requiring that the events {X = x} and A are independent, for any choice x.

• Consider P(A) > 0

By the definition of the conditional PMF  $p_{X|A}(x) = P(X = x \text{ and } A)/P(A)$ 

Independence is the same as the condition  $p_{X|A}(x) = p_X(x)$  for all x

- Consider two independent tosses of a fair coin.
  - □ X: the number of heads
  - □ A: the number of heads is even
- The PMF of X

$$p_X(x) = \begin{cases} 1/4 & \text{if } x = 0\\ 1/2 & \text{if } x = 1\\ 1/4 & \text{if } x = 2 \end{cases}$$

• We know  $P(A) = \frac{1}{2}$ 

The conditional PMF

$$p_{X|A}(x) = \begin{cases} 1/2 & \text{if } x = 0\\ 0 & \text{if } x = 1\\ 1/2 & \text{if } x = 2 \end{cases}$$

• The PMFs  $p_X$  and  $p_{X|A}$  are different  $\Rightarrow X$  and A are not independent

- The notion of independence of two random variables is similar.
- Two random variables X and Y are independent if

 $p_{X,Y}(x,y) = p_X(x)p_Y(y)$  for all x, y

Same as requiring that the two events {X = x} and {Y = y} be independent for every x and y.

- By the formula  $p_{X,Y}(x,y) = p_{X|Y}(x|y)p_Y(y)$
- Independence is equivalent to the condition  $p_{X|Y}(x|y) = p_X(x)$

for all y with  $p_Y(y) > 0$  and all x.

Independence means that the experimental value of Y tells us nothing about the value of X.

- X and Y are conditionally independent, if given a positive probability event A P(X = x, Y = y|A) = P(X = x|A)P(Y = y|A)
- Using this chapter's notation  $p_{X,Y|A}(x,y) = p_{X|A}(x)p_{Y|A}(y)$
- Or equivalently,

$$p_{X|Y,A}(x|y) = p_{X|A}(x)$$

for all x, y such that  $p_{Y|A}(y) > 0$ .

If X and Y are independent random variables, then

$$\mathbf{E}[XY] = \mathbf{E}[X] \cdot \mathbf{E}[Y]$$
  
Shown by the following calculation  
$$\mathbf{E}[XY] = \sum_{x} \sum_{y} xy \cdot p_{X,Y}(x, y)$$
$$= \sum_{x} \sum_{y} xy \cdot p_{X}(x)p_{Y}(y)$$
$$= \sum_{x} xp_{X}(x) \cdot \sum_{y} yp_{Y}(y)$$
$$= \mathbf{E}[X] \cdot \mathbf{E}[Y]$$

- Conditional independence may not imply unconditional independence.
- X and Y are not independent

□ 
$$p_{X|Y}(1|1) = P(X = 1|Y = 1)$$
  
= 0 ≠  $P(X = 1) = p_X(1)$ 

Condition on

 $A = \{X \le 2, Y \ge 3\}$ 

They are independent

y ,					
4	1/20	2/20	2/20	0	
3	2/20	4/20	1/20	2/20	
2	0	1/20	3/20	1/20	
1	0	1/20	0	0	
	1	2	3	4	x

- A very similar calculation shows that if X and Y are independent, then so are g(X) and h(Y) for any functions g and h.
- $\mathbf{E}[g(X)h(Y)] = \mathbf{E}[g(X)]\mathbf{E}[h(Y)]$
- Next, we consider variance of sum of independent random variables.

- Consider Z = X + Y, where X and Y are independent.
- $\operatorname{Var}(Z) = \operatorname{E}[(X + Y \operatorname{E}[X + Y])^2]$   $= \operatorname{E}[(X + Y - \operatorname{E}[X] - \operatorname{E}[Y])^2]$   $= \operatorname{E}\left[\left((X - \operatorname{E}[X]) + (Y - \operatorname{E}[Y])\right)^2\right]$   $= \operatorname{E}[(X - \operatorname{E}[X])^2] + \operatorname{E}[(Y - \operatorname{E}[Y])^2]$  $+ 2\operatorname{E}[(X - \operatorname{E}[X])(Y - \operatorname{E}[Y])]$

- Now we compute  $\mathbf{E}[(X \mathbf{E}[X])(Y \mathbf{E}[Y])]$ .
- Since X and Y are independent, so are  $X \mathbf{E}[X]$  and  $Y \mathbf{E}[Y]$ .
  - As they are two functions of X and Y, respectively.
- Thus E[(X E[X])(Y E[Y])]
  - $= \mathbf{E}[(X \mathbf{E}[X])] \cdot \mathbf{E}[(Y \mathbf{E}[Y])]$

 $= 0 \cdot 0 = 0$ 

• So  $\operatorname{Var}(Z) = \operatorname{E}[(X - \operatorname{E}[X])^2] + \operatorname{E}[(Y - \operatorname{E}[Y])^2]$ =  $\operatorname{Var}[X] + \operatorname{Var}[Y]$  Summary of independent r.v.'s

• X is independent of the event A if  $p_{X|A}(x) = p_X(x)$ 

that is, if for all x, the events  $\{X = x\}$  and A are independent.

• X and Y are independent if for all possible pairs (x, y), the events  $\{X = x\}$  and  $\{Y = y\}$  are independent

$$p_{X,Y}(x,y) = p_X(x)p_Y(y)$$

Summary of Facts About Independent Random Variables

- If X and Y are independent random variables, then
  - 1.  $\mathbf{E}[XY] = \mathbf{E}[X]\mathbf{E}[Y]$
  - 2.  $\mathbf{E}[g(X)h(Y)] = \mathbf{E}[g(X)]\mathbf{E}[h(Y)]$ , for any functions g and h.
  - 3.  $\operatorname{Var}[X + Y] = \operatorname{Var}[X] + \operatorname{Var}[Y]$

#### Independence of Several Random Variables

- All previous results have natural extensions to more than two random variables.
- Example: Random variables X, Y, and Z are independent if

$$p_{X,Y,Z}(x,y,z) = p_X(x)p_Y(y)p_Z(z)$$

• Example: If  $X_1, X_2, ..., X_n$  are independent random variables, then

$$Var(X_1 + X_2 + \dots + X_n)$$
  
= Var(X\_1) + Var(X\_2) + \dots + Var(X\_n)

# Variance of the Binomial

- Consider n independent coin tosses
  - $\square P(H) = p$
  - $X_i$ : Bernoulli random variable for *i*th toss

#### Its PMF

# $p_{X_i}(x) = \begin{cases} 1 & i \text{th toss comes up a head} \\ 0 & \text{otherwise} \end{cases}$

# Variance of the Binomial

- Let  $X = X_1 + X_2 + \dots + X_n$  be a binomial random variable.
- By the independence of the coin tosses  $\operatorname{Var}(X) = \sum_{i=1}^{n} \operatorname{Var}(X_i) = np(1-p)$

#### Mean and Variance of the Sample Mean

Estimate the approval rating of a president *C*.
 Ask *n* persons randomly from the voters
 X<sub>i</sub> response of the *i*th person

• 
$$X_i = \begin{cases} 1 & i \text{th person approves } C \\ 0 & i \text{th person disapproves } C \end{cases}$$

Mean and Variance of the Sample Mean

- Model X<sub>1</sub>, X<sub>2</sub>, ..., X<sub>n</sub> as independent Bernoulli random variables
  - mean p
  - variance p(1-p)
  - The sample mean  $S_n = \frac{X_1 + X_2 + \dots + X_n}{n}$

#### Mean and Variance of the Sample Mean

- S<sub>n</sub> is the approval rating of C within our n-person sample.
- Using the linearity of  $S_n$  as a function of the  $X_i$  $\mathbf{E}[S_n] = \sum_{i=1}^n \frac{1}{n} \mathbf{E}[X_i] = \frac{1}{n} \sum_{i=1}^n p = p$

and

$$\operatorname{Var}(S_n) = \sum_{i=1}^n \frac{1}{n^2} \operatorname{Var}(X_i) = \frac{p(1-p)}{n}$$