Introduction to Queueing Networks

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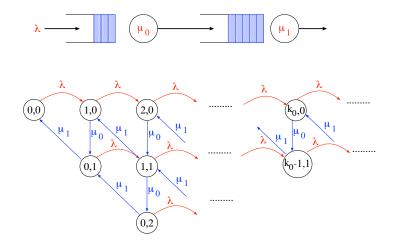
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Introduction Illu

Illustration

Example: Network of Queues



• for
$$k_0 > 0, k_1 > 0$$
:
 $(\mu_0 + \mu_1 + \lambda)p(k_0, k_1) = \mu_0 p(k_0 + 1, k_1 - 1) + \mu_1 p(k_0, k_1 + 1) + \lambda p(k_0 - 1, k_1)$
• for $k_0 > 0, k_1 = 0$:
 $(\mu_0 + \lambda)p(k_0, 0) = \mu_1 p(k_0, 1) + \lambda p(k_0 - 1, 0)$
• for $k_0 = 0, k_1 > 0$:
 $(\mu_1 + \lambda)p(0, k_1) = \mu_0 p(1, k_1 - 1) + \mu_1 p(0, k_1 + 1)$
• for $k_0 = 0, k_1 = 0$:
 $\lambda p(0, 0) = \mu_1 p(0, 1)$

Normalization :

1

$$\sum_{k_0=0}^{\infty} \sum_{k_1=0}^{\infty} p(k_0, k_1) = 1$$

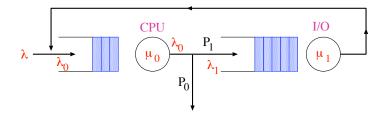
$$p(k_0, k_1) = (1 - \rho_0) \rho_0^{k_0} (1 - \rho_1) \rho_1^{k_1} \text{ for } k_0, k_1 = 0, 1, \cdots$$

where
$$ho_0 = rac{\lambda}{\mu_0}$$
; $ho_1 = rac{\lambda}{\mu_1}$;
 $P[N_0 = k_0] = (1 -
ho_0)
ho_0^{k_0}$; $P[N_1 = k_1] = (1 -
ho_1)
ho_1^{k_1}$

Example

Open Queueing Network [Jackson 57]

It allows "feedback" and "product-form" can still be maintained.



$$p(k_0, k_1) = (1 - \rho_0) \rho_0^{k_0} (1 - \rho_1) \rho_1^{k_1}$$

$$\lambda_0 = \lambda + \lambda_1$$

$$\lambda_1 = \lambda_0 p_1$$

Therefore,

$$\lambda_0 = \frac{\lambda}{1 - p_1}$$
$$\lambda_1 = \frac{\lambda p_1}{1 - p_1}$$
$$\rho_0 = \frac{\lambda_0}{\mu_0} = \frac{\lambda}{(1 - p_1)\mu_0} = \frac{\lambda}{p_0\mu_0}$$
$$\rho_1 = \frac{\lambda_1}{\mu_1} = \frac{\lambda p_1}{p_0\mu_1}$$

What is the average response time of a job?

By Little's formula, $T = \frac{\bar{N}}{\lambda}$ where \bar{N} is the average no. of jobs in the "system".

$$\bar{N} = \bar{N}_0 + \bar{N}_1 = \sum_{k_0=0}^{\infty} \sum_{k_1=0}^{\infty} (k_0 + k_1) \rho(k_0, k_1)$$

$$= \sum_{k_0=0}^{\infty} \sum_{k_1=0}^{\infty} k_0 \rho(k_0, k_1) + \sum_{k_0=0}^{\infty} \sum_{k_1=0}^{\infty} k_1 \rho(k_0, k_1)$$

$$= \sum_{k_0=0}^{\infty} k_0 (1 - \rho_0) \rho_0^{k_0} \sum_{k_1=0}^{\infty} (1 - \rho_1) \rho_1^{k_1}$$

$$+ \sum_{k_0=0}^{\infty} (1 - \rho_0) \rho_0^{k_0} \sum_{k_1=0}^{\infty} k_1 (1 - \rho_1) \rho_1^{k_1}$$

$$\bar{N} = \frac{\rho_0}{1 - \rho_0} + \frac{\rho_1}{1 - \rho_1} ; \quad E[T] = \frac{\bar{N}}{\lambda} = \left[\frac{\rho_0}{1 - \rho_0} + \frac{\rho_1}{1 - \rho_1}\right] \left[\frac{1}{\lambda}\right]$$

- Types of service centers
 - FCFS and service time is exponentially distributed.
 - Processor sharing (PS)
 - Last come first serve pre-emptive resume (LCFS-PR)
 - Infinite server (IS) or delay nodes
- We also allow a state dependent service rate (μ_i(n) = service rate at the ith node where there is n customer).
 - Single server fixed rate (SSFR) where $\mu_i(n) = u_i$
 - Infinite server (IS) , $\mu_i(n) = n\mu_i$
 - Queue length dependent (QLD) with service rate $\mu_i(n)$

Jackson Network

- A queueing network with *M* nodes (labeled $i = 1, 2, \dots, M$) s.t.
- Node i is QLD with rate $\mu_i(n)$ when it has n customers.
- A customer completing service at a node makes a probabilistic choice of either leaving the network or entering another node, independent of past history.
- The network is open and any external arrivals to node *i* is from a **Poisson** stream.

Jackson Network : Continue

State space
$$S = \{(n_1, n_2, \cdots, n_M) \mid n_i \ge 0\}$$

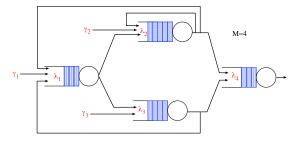
Routing probability matrix $Q = (q_{ij} | i, j = 1, 2, \dots M)$

$$q_{i0}=1-\sum_{j=1}^M q_{ij}$$

 Let γ be the <u>"TOTAL" external arrival rate</u> to the open queueing network, the rate to node i is γ_i = γq_{0i} for i = 1, 2, · · · M. So

$$\gamma = \sum_{i=1}^{M} \gamma_i \qquad (\sum q_{0i} = 1)$$

 All nodes in the Jackson network are QLD with exponential service time. Pictorially we can have:



• Let λ_i = mean arrival rate to node i, ($i = 1, 2, \dots, M$)

$$\lambda_i = \gamma_i + \sum_{j=1}^M \lambda_j q_{ji}$$
 there is **unique solution** to $\{\lambda_i\}$

• Example :

$$\lambda_{1} = \gamma_{1} + \lambda_{2}q_{21} + \lambda_{3}q_{31}$$
$$\lambda_{2} = \gamma_{2} + \lambda_{1}q_{12} + \lambda_{2}q_{22}$$
$$\lambda_{3} = \gamma_{3} + \lambda_{1}q_{13}$$
$$\lambda_{4} = \lambda_{2}q_{24} + \lambda_{3}q_{34}$$

Jackson's Theorem: For a Jackson Network in steady state with arrival rate λ_i to node *i*,

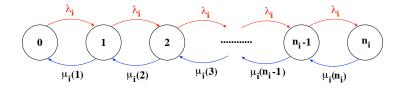
- The no. of customer at any node is independent of the number of customers at every other node.
- Node *i* behaves "stochastically" <u>as if it were</u> subjected to Poisson arrival rates λ_i
- Let $\pi(\vec{n}) = \operatorname{Prob}[(n_1, n_2, \cdots n_M)]$ where $n_i \ge 0$, for **SSFR**:

$$\pi(\vec{n}) = \prod_{i=1}^{M} \left(1 - \frac{\lambda_i}{\mu_i}\right) \left(\frac{\lambda_i}{\mu_i}\right)^{n_i}$$

• For **QLD**, M/M/c queues, define $u_i(r) = \mu_i \min(r, c_i)$ for $r \ge 0$, i = 1, ..., M, and $\rho_i = \lambda_i/\mu_i$ for i = 1, ...M.

$$\pi(\vec{n}) = \prod_{i=1}^{M} C_{i} \left(\frac{\lambda_{i}^{n_{i}}}{\prod_{r=1}^{n_{i}} \mu_{i}(r)} \right) \quad ; \quad C_{i} = \left[\sum_{r=0}^{c_{i}-1} \frac{\rho_{i}^{r}}{r!} + \left(\frac{\rho_{i}^{c_{i}}}{c_{i}!} \right) \left(\frac{1}{1 - \rho_{i}/c_{i}} \right) \right]^{-1}$$

The *i*th node:



How about the normalization constant C_i ?

$$\pi(\vec{n}) = \pi(n_1, n_2, \cdots, n_M) = \prod_{i=1}^M \pi_i(n_i) = \prod_{i=1}^M C_i \left[\frac{\lambda_i^{n_i}}{\prod_{j=1}^{n_i} \mu_j(j)} \right]$$

where C_i can be found by $\sum_{n_i=0}^{\infty} C_i \left[\frac{\lambda_i^{n_i}}{\prod_{j=1}^{n_j} \mu_i(j)} \right] = 1$

- So , what do we have?
 - extremely powerful modeling tool to model a very large class of system.
 - efficient solution
 - we can compute mean queue length, utilization and throughout
 - mean response time.

Comment

- The arrival to each node, in general (unless it's only feed forward), is NOT a Poisson process.
- **2** How can we compute the \overline{T} and each node utilization?
- Optimal allocation : assume that the open network of SSFR nodes with arrival rate λ_i and μ_i for each node ($i = 1, 2, \dots M$)

$$\begin{array}{rcl} \text{Min} & \bar{N} & = & \displaystyle\sum_{i=1}^{M} \frac{\frac{\lambda_i}{\mu_i}}{1 - \frac{\lambda_i}{\mu_i}} \\ \text{s.t.} & \displaystyle\sum_{i=1}^{M} \mu_i & = & C = \text{constant} \end{array}$$

Application : Network, distributed system

Example 1

Consider a switching facility that transmits messages to a required destination. A NACK is sent by the destination when a packet has not been properly received. If so, the packet in error is retransmitted as soon as the NACK is received.

Assume the time to send a message and the time to receive a NACK are both exponentially distributed with parameter μ . Assume that packets arrive at the switch according to a Poisson process with rate λ^0 . Let p, 0 , be the probability that a message is received correctly. Derive mean response time <math>T.

We can model it as a Jackson network of one node with feedback, where $c_i = 1$ (SSFR), $p_{10} = p$ and $p_{11} = 1 - p$. Let $\pi(n)$, the probability of having *n* packets, is:

$$\lambda = \lambda^{0} + \lambda(1 - p)$$

$$\pi(n) = \left(1 - \frac{\lambda}{\mu}\right) \left(\frac{\lambda}{\mu}\right)^{n} \quad n \ge 0$$

We have $\lambda = \lambda^0 / p$ and

$$\pi(n) = \left(1 - \frac{\lambda^0}{p\mu}\right) \left(\frac{\lambda^0}{p\mu}\right)^n \quad n \ge 0$$

Let N and T be the mean number of packet and mean response time.

$$egin{array}{rcl} N & = & \displaystylerac{\lambda^0}{p\mu-\lambda^0} \ T & = & \displaystylerac{1}{p\mu-\lambda^0} \end{array}$$

Example 2

Similar to last example but now the switching facility is composed of *K* nodes in series, each model as M/M/1 queue with switching rate μ . What is the response time *T*?

• We have $\lambda_i^0 = 0$ for i = 2, ..., K (no external arrival to nodes 2, ..., K), $\mu_i = \mu$ for i = 1, 2, ..., K, $p_{i,i+1} = 1$ for i = 1, ..., K - 1, and $p_{K,0} = p$ and $p_{K,1} = 1 - p$. • $\lambda_1 = \lambda^0 + (1 - p)\lambda_K$, $\lambda_i = \lambda_{i-1}$ for i = 2, ..., K. So

$$\lambda_i = \lambda^0 / \boldsymbol{\rho} \qquad \forall i = 1, ..., K.$$

By the Jackson's theorem, we have

$$\boldsymbol{\pi}(\vec{n}) = \left(\frac{\boldsymbol{p}\mu - \lambda^{0}}{\boldsymbol{p}\mu}\right)^{K} \left(\frac{\lambda^{0}}{\boldsymbol{p}\mu}\right)^{n_{1}+\dots+n_{K}} \quad \forall \vec{n} = (n_{1}, n_{2}, \dots, n_{K}) \in \mathbb{N}^{K}$$

provided that $\lambda^0 < p\mu$. Let $E[N_i]$ be the average number of packets in queue *i*:

$$E[N_i] = rac{\lambda^0}{p\mu - \lambda^0}$$
 $i = 1, \dots, K.$

Let E[T] be the average response time:

$$E[T] = \sum_{i=1}^{K} E[N_i] = K\left(\frac{1}{p\mu - \lambda^0}\right).$$

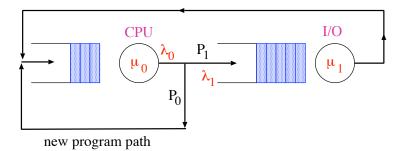
Example 3: open central server network

- A computer system with one CPU and two I/O devices. New jobs enter the system, wait for CPU resource, then possibly an I/O requests. When a job finishes its I/O, it may return for more CPU resource. Eventually a job completes and leave the system.
- It means that K = 3 (three nodes). $\lambda_i^0 = 0$ for i = 2, 3, $p_{2,1} = p_{3,1} = 1$, while $p_{1,0} > 0$.
- The traffic equations are: $\lambda_1 = \lambda_1^0 + \lambda_2 + \lambda_3$, $\lambda_2 = \lambda_1 p_{1,2}$, $\lambda_3 = \lambda_1 p_{1,3}$. Solving, we have: $\lambda_1 = \lambda_1^0 / p_{1,0}$, $\lambda_i = \lambda_1^0 p_{1,i} / p_{1,0}$ for i = 2, 3. Thus,

$$\pi(\vec{n}) = \left(1 - \frac{\lambda_1^0}{\mu_1 \rho_{1,0}}\right) \left(\frac{\lambda_1^0}{\mu_1 \rho_{1,0}}\right)^{n_1} \prod_{i=2}^3 \left(1 - \frac{\lambda_1^0 \rho_{1,i}}{\mu_i \rho_{1,0}}\right) \left(\frac{\lambda_1^0 \rho_{1,i}}{\mu_i \rho_{1,0}}\right)^{n_i} \quad \vec{n} \in \mathbb{N}^3$$
$$E[T] = \frac{1}{\mu_1 \rho_{1,0} - \lambda_1^0} + \sum_{i=2}^3 \frac{\rho_{1,i}}{\mu_i \rho_{1,0} - \lambda_1^0 \rho_{1,i}}$$

Example of Queueing Network

• We fixed the total number of jobs be *n* in the system, where *n* is also called the "*degree of multiprogramming*".



• State representation (k_0, k_1) . Now with the constant that $k_0 + k_1 = n$

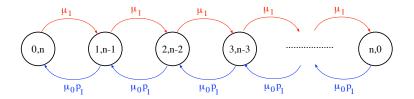
• Unlike the open queueing network, the state space is **finite**. The flow balance equations are:

$$(\mu_1 + \mu_0 p_1) p(k_0, k_1) = \mu_0 p_1 p(k_0 + 1, k_1 - 1) + \mu_1 p(k_0 - 1, k_1 + 1)$$

$$\mu_1 p(0, n) = \mu_0 p_1 p(1, n - 1)$$

$$\mu_0 p_1 p(n, 0) = \mu_1 p(n - 1, 1)$$

If we draw the state transition diagram, we have:



- Using "traditional method", we know how to find the solution of p(k₀, k₁)
- New approach: Let $\rho_0 = \frac{a}{\mu_0}$; $\rho_1 = \frac{a\rho_1}{\mu_1}$, by substituting to the flow equations

$$p(k_0, k_1) = \frac{1}{C(n)} \rho_0^{k_0} \rho_1^{k_1} \qquad k_0, k_1 \ge 0$$

• The normalization factor *C*(*n*) is chosen s.t

$$\sum_{k_0+k_1=n\,\&\,k_0,k_1\geq 0} p(k_0,k_1) = 1$$

- The choice of *a* (where $\rho_0 = \frac{a}{\mu_0}$, $\rho_1 = \frac{a\rho_1}{\mu_1}$) is very *arbitrary* in that the value of $p(k_0, k_1)$ will not change with a.
- If we define λ₀ = a and λ₁ = ap₁, then (λ₀, λ₁) as the *relative throughput* of the corresponding nodes.
- We can choose, for example, a = 1 or $a = \mu_0$. Assume we choose $a = \mu_0$, then $\rho_0 = 1$, $\rho_1 = \frac{\mu_0 \rho}{\mu_1}$. Since

$$p(k_0, k_1) = \frac{1}{C(n)} \rho_0^{k_0} \rho_1^{k_1} = \frac{1}{C(n)} \rho_1^{k_1}$$

$$1 = \frac{1}{C(n)} \sum_{k_1=0}^n \rho_1^{k_1} = \frac{1}{C(n)} \left[\frac{1 - \rho_1^{n+1}}{1 - \rho_1} \right]$$

$$C(n) = \begin{cases} \frac{1-\rho_1^{n+1}}{1-\rho_1} & \text{where } \rho_1 \neq 1\\ n+1 & \text{where } \rho_1 = 1(\text{L'Hospital Rule}) \end{cases}$$

If we choose a = 1, then $\rho_0 = \frac{1}{\mu_0}$, $\rho_1 = \frac{p_1}{\mu_1}$

$$p(k_{0}, k_{1}) = \frac{1}{C(n)} \rho_{0}^{k_{0}} \rho_{1}^{k_{1}} = \frac{1}{C(n)} \left(\frac{1}{\mu_{0}}\right)^{k_{0}} \left(\frac{p_{1}}{\mu_{1}}\right)^{k_{1}}, \text{ summing all } k_{0}, k_{1}$$

$$1 = \sum_{k_{0}=0}^{n} \sum_{k_{1}=n-k_{0}}^{n-k_{0}} \frac{1}{C(n)} \left(\frac{1}{\mu_{0}}\right)^{k_{0}} \left(\frac{p_{1}}{\mu_{1}}\right)^{k_{1}}$$

$$= \sum_{k_{0}=0}^{n} \frac{1}{C(n)} \left(\frac{1}{\mu_{0}}\right)^{k_{0}} \left(\frac{p_{1}}{\mu_{1}}\right)^{n-k_{0}}$$

$$C(n) = \sum_{k_{0}=0}^{n} \left(\frac{1}{\mu_{0}}\right)^{k_{0}} \left(\frac{p_{1}}{\mu_{1}}\right)^{n-k_{0}} = \left(\frac{p_{1}}{\mu_{1}}\right)^{n} \sum_{k_{0}=0}^{n} \left(\frac{1}{\mu_{0}}\right)^{k_{0}} \left(\frac{p_{1}}{\mu_{1}}\right)^{-k_{0}}$$

$$= \left(\frac{p_{1}}{\mu_{1}}\right)^{n} \sum_{k_{0}=0}^{n} \left(\frac{\mu_{1}}{\mu_{1}\mu_{0}}\right)^{n+1}$$

$$C(n) = \left(\frac{p_{1}}{\mu_{1}}\right)^{n} \left[\frac{1-\left(\frac{\mu_{1}}{\mu_{1}\mu_{0}}\right)^{n+1}}{1-\frac{\mu_{1}}{\mu_{1}\mu_{0}}}\right]$$

$$\rightarrow p(k_0, k_1) = \frac{1}{C(n)} \left(\frac{1}{\mu_0}\right)^{k_0} \left(\frac{p_1}{\mu_1}\right)^{k_1} \\ = \frac{1}{C(n)} \left(\frac{1}{\mu_0}\right)^{k_0} \left(\frac{p_1}{\mu_1}\right)^{n-k_0} = \frac{1}{C(n)} \left(\frac{p_1}{\mu_0}\right)^n \left(\frac{\mu_1}{p_1\mu_0}\right)^{k_0} \\ p(k_0, k_1) = \left(\frac{\mu_1}{p_1}\right)^n \left[\frac{1 - \frac{\mu_1}{p_1\mu_0}}{1 - (\frac{\mu_1}{p_1\mu_0})^{n+1}}\right] \left(\frac{p_1}{\mu_1}\right)^n \left(\frac{\mu_1}{p_1\mu_0}\right)^{k_0} \\ = (p_1)^{-k_0} \left[\frac{1 - p_1^{-1}}{1 - p_1^{-(n+1)}}\right] \\ = (p_1)^{-(n-k_1)} \left[p_1^n - \frac{1 - p_1}{1 - p_1^{n+1}}\right] \\ = (p_1)^{k_1} \left(\frac{1 - p_1}{1 - p_1^{n+1}}\right)$$

• The CPU utilization is

CPU utilization = Prob[CPU is busy]
=
$$1 - P(0, n) = 1 - \frac{\rho_1^n}{C(n)}$$

= $\begin{cases} \frac{\rho_1 - \rho_1^{n+1}}{1 - \rho_1^{n+1}} & \rho_1 \neq 1\\ \frac{n}{n+1} & \rho_1 = 1 \end{cases}$

Average throughput is

$$E[T] = P[CPU ext{ is busy}]\mu_0 p_0$$

Gordon - Newell network (1967)

• A Gordon-Newell network has M nodes ($i = 1, 2, \dots M$) s.t.

- Node *i* is QLD with rate $\mu_i(n)$ when there is *n* customers.
- a customer completing service at a node chooses a node to enter next probabilistically, independent of past history
- The network is **CLOSED** and has a fixed population *K*
- State space: $S = \{(n_1, n_2, \dots n_M) \mid n_i \ge \phi, \sum_{i=1}^M n_i = K\}$

$$|S| = \begin{pmatrix} K + M - 1 \\ M - 1 \end{pmatrix} \rightarrow$$
VERY LARGE NUMBER

If M = 5, K = 10, |S| = 1,001. If M = 10, K = 35, |S| = 52,451,256. • Since there is **no** external arrival, $\gamma = \phi$.

• Routing probabilities *q_{ij}* satisfy:

$$\sum_{j=1}^{M} q_{ij} = 1$$

• traffic equations are:

$$\lambda_i = \sum_{j=1}^M \lambda_j q_{ji}$$
 $i = 1, 2, \cdots, M$

- The number of solutions {λ_i} that satisfy the traffic equations is infinite.
- All solutions differ by a multiplicative factor C
- Let $(e_1, e_2, \dots e_M)$ be any non-zero solution, that is $e_i = C\lambda_i$ (visit rate). Define: $x_i = \frac{e_i}{u_i}$
- (*e*₁, *e*₂, ··· *e*_M) is chosen by fixing one component to a convenient value, such as *e*₁ = 1

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Gordon Newell Theorem

$$\pi(n_1, n_2, \cdots n_M) = \frac{1}{G} \prod_{i=1}^M x_i(n_i)$$
where $x_i(n_i) = \left[\frac{e_i^{n_i}}{\prod_{j=1}^{n_i} \mu_i(j)}\right]$ and $\sum_{i=1}^M n_i = K$
Therefore,

$$G = \sum_{\vec{n} \in S} \prod_{i=1}^{m} x_i(n_i)$$

Computation of G (assume SSFR)

Define
$$S(m, n) = \{(n_1, \dots n_m) \mid n_i \ge 0, \sum_{i=1}^m n_i = n\}$$

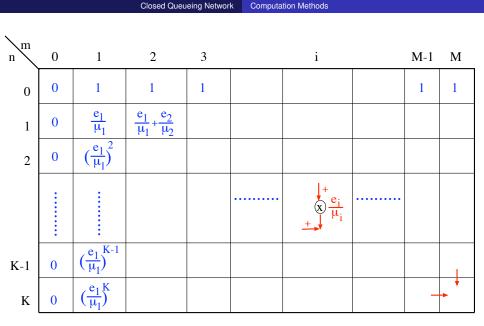
$$G(m,n) = \sum_{\vec{n} \in S(m,n)} \prod_{i=1}^{m} x_i(n_i)$$
 where $x_i = \frac{e_i}{\mu_i}$

$$G(m, n) = \sum_{\substack{\vec{n} \in S(m, n); \\ n_m = 0}} \prod_{i=1}^m x_i^{n_i} + \sum_{\substack{\vec{n} \in S(m, n); \\ n_m > \phi}} \prod_{i=1}^m x_i^{n_i}$$
$$= \sum_{\vec{n} \in S(m-1, n)} \prod_{i=1}^{m-1} x_i(n_i) + x_m \sum_{\substack{\vec{n} \in S(m, n); \\ k_i = n_i(i \neq m); \\ k_m = n_m - 1}} \prod_{i=1}^m x_i(k_i)$$

$$G(m,n) = G(m-1,n) + x_m G(m,n-1) \quad m,n > \phi$$

In summary, for SSFR queueing network, we have

$$\begin{array}{rcl} G(m,n) & = & G(m-1,n) + \frac{e_m}{\mu_m} G(m,n-1) & m,n > 0 \\ G(m,0) & = & 1 & m > 0 \\ G(0,n) & = & 0 & n \ge 0 \end{array}$$



Performance Measures

• Idle Probability for node *i*:

$$P(N_{M} = 0) = \frac{1}{G(M, K)} \sum_{\substack{\vec{n} \in S(M, K): \\ n_{M} = 0}} \prod_{i=1}^{M-1} \left(\frac{e_{i}}{\mu_{i}}\right)^{n_{i}} = \frac{G(M-1, K)}{G(M, K)}$$

Another way to express it:

$$P(N_i = 0) = \frac{1}{G(M, K)} \sum_{\substack{\vec{n} \in S(\mu, k): \\ n_i = 0}} \prod_{j=1}^{i-1} \left(\frac{e_i}{\mu_j}\right)^{n_j} \prod_{k=i+1}^M \left(\frac{e_k}{\mu_k}\right)^{n_k}$$
$$= \frac{G(M \setminus i, K)}{G(M, K)}$$

• Utilization of node *i* or (U_i) :

$$U_i = 1 - P[N_i = 0] = 1 - \frac{G(M \setminus i, K)}{G(M, K)}$$

· - - · ·

• How about $P[N_i \ge k]$:

$$P[N_i \ge k] = \frac{1}{G(M, K)} \sum_{\substack{\vec{n} \in S(M, K): \\ n_i \ge k}} \prod_{j=1}^M \left(\frac{\mathbf{e}_j}{\mu_j}\right)^{n_j}$$
$$= \frac{1}{G(M, K)} \left(\frac{\mathbf{e}_i}{\mu_i}\right)^k \sum_{\substack{m_j = n_j (i \neq i): \\ m_j = n_j - k: \\ \vec{n} \in S(M, K): \\ n_j \ge k}} \prod_{j=1}^M \left(\frac{\mathbf{e}_j}{\mu_j}\right)^{m_j} = \left(\frac{\mathbf{e}_i}{\mu_i}\right)^k \frac{G(M, K - k)}{G(M, K)}$$

• For $P[N_i \ge 1]$:

$$P[N_i \ge 1] = \mu_i = \left(\frac{e_i}{\mu_i}\right) \frac{G(M, K-1)}{G(M, K)}$$

(more useful)

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The **throughout** of node *i* is

$$\lambda_i = \mu_i U_i = \mu_i \left(\frac{e_i}{\mu_i}\right) \frac{G(M, K-1)}{G(M, K)} = e_i \frac{G(M, K-1)}{G(M, K)}$$

 $Prob[n_i = k] = \pi_i(k) = P[N_i > k] - P[N_i > k + 1]$ $\pi_i(k) = \frac{1}{G(M,K)} \sum_{\vec{n} \in S(M,K)} \prod_{i=1}^M \left(\frac{\mathbf{e}_i}{\mu_i}\right)^{n_i}$ $= \frac{\left(\frac{e_i}{\mu_i}\right)^{\kappa}}{G(M,K)} \sum_{\vec{n} \in S(M,K-k); j=1} \prod_{j=1}^{M} \left(\frac{e_j}{\mu_j}\right)^{n_j} = \left(\frac{e_i}{\mu_i}\right)^{\kappa} \frac{G(M \setminus i, K-k)}{G(M,K)}$ $\pi_i(k) = P[N_i > k] - P[N_i > k + 1]$ $= \left(\frac{e_i}{\mu_i}\right)^{\kappa} \left| \frac{G(M, K-k) - \frac{e_i}{\mu_i}G(M, K-k-1)}{G(M, K)} \right|$

Expected no. of customer in node $i = L_i(K)$

$$L_{i}(K) = \sum_{j=1}^{K} j P[N_{i} = j] = \sum_{j=1}^{K} P[N_{i} = j] \sum_{k=1}^{K} I_{(k \le j)}$$

$$= \sum_{k=1}^{K} \sum_{j=k}^{K} P[N_{i} = j] = \sum_{k=1}^{K} P[N_{i} \ge k]$$

$$L_{i}(K) = \frac{1}{G(M, K)} \sum_{k=1}^{K} \left(\frac{e_{i}}{\mu_{i}}\right)^{k} G(M, K - k) \quad i = 1, 2, \cdots, M$$

Derivation of above expression:

$$L_{i}(K) = P[N_{i} = 1] + P[N_{i} = 2] + P[N_{i} = 2] + P[N_{i} = 3] + P[N_{i} = 3] + P[N_{i} = 3] + P[N_{i} = 3] + \cdots + P[N_{i} = K] + P[N_{i} = K] + P[N_{i} = K] + \cdots + P[N_{i} = K]$$

Garden Newell Convolution Algorithm

- Assume M QLD nodes and K customers
- Define

$$f_i(Z) = \sum_{k=0}^{\infty} X_i(k) Z^k = 1 + X_i(1) Z + X_i(2) Z^2 + X_i(3) Z^3 + \cdots$$

$$= 1 + \left[\frac{e_i}{\mu_i(1)}\right] Z + \left[\frac{e_i^2}{\mu_i(1)\mu_i(2)}\right] Z^2 + \left[\frac{e_i^3}{\mu_i(1)\mu_i(2)\mu_i(3)}\right] Z^3 + \cdots$$

$$f(Z) = f_1(Z) f_2(Z) \dots f_M(Z)$$

• The coefficient of Z^k in $f(Z) \rightarrow$ the sum of products of the form $X_1(n_1)X_2(n_2)\ldots X_M(n_M)$ such that $\sum_i n_i = k$

Thus

$$f(Z) = 1 + G(M, 1)Z + G(M, 2)Z^2 + G(M, 3)Z^3 + \dots + G(M, k)Z^k + \dots$$

 *IDEA: build up f(Z) from partial products g_i(Z) so that G(i, k) is the coefficient of Z^k in g_i(Z)

$$g_1(Z) = f_1(Z) g_i(Z) = g_{i-1}(Z)f_i(Z) \quad i = 2, 3, \cdots, M$$

Therefore,

$$G(1,k) = \text{coefficient of the } Z^k \text{ term in } g_1(Z)$$

$$G(1,k) = X_1(k) = \frac{e_1^k}{\prod_{i=1}^k \mu_1(i)}$$

$$G(i,k) = \sum_{j=0}^k G(i-1,j)x_i(k-j) \to (\text{convolution!})$$

$$G(2,k) = \sum_{j=0}^{k} G(1,j) x_2(k-j)$$

= $G(1,0) x_2(k) + G(1,1) x_2(k-1) + \dots + G(1,k) x_2(\phi)$
= $\frac{e_2^k}{\prod_{j=1}^k x_2(j)} + \frac{e_1}{\mu_1(1)} \frac{e_2^{k-1}}{\prod_{j=1}^{k-1} \mu_2(j)} + \dots$

$k \setminus \mu$	1	2	• • •
0	1	1	• • •
1	$\frac{\frac{e_1}{\mu_1(1)}}{e_1^2}$		
2	$rac{e_1^2}{\mu_1(1)\mu_1(2)}$		
3	$\frac{\overline{\mu_1(1)\mu_1(2)}}{e_1^3}$		
К	$rac{e_1^k}{\mu_1(1)\mu_1(k)}$		

Performance Measure

• Prob[node *i* has *k* customers] = $\pi_i(k)$:

$$\pi_i(k) = \frac{1}{G(M, K)} \sum_{\substack{\vec{n} \in S(\mu, k):\\ n_i = k}} X_1(n_1) X_2(n_2) \dots X_M(n_M)$$

$$\pi_i(k) = \frac{x_i(k)}{G(M,K)}G(M \setminus i, K-k)$$

• Expected number of customers in node *i*

$$E[L_i] = \sum_{k=0}^{K} k \pi_i(k) \Longrightarrow$$
 very involve

• Prob[node *i* ≥ 1 customer] = ?

• Utilization of a node, in general, is very involved.

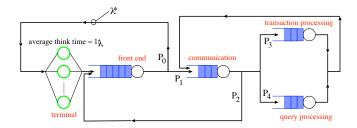
$$U_i(K) = Prob[n_i \ge 1] = \sum_{j=1}^{K} Prob[n_i = j]$$

For a node that is SSFR, we have

$$U_i(K) = \left(\frac{e_i}{\mu_i}\right) \frac{G(M, K-1)}{G(M, K)}$$

• This holds even the other nodes are QLD servers!!

Example :



 $\begin{aligned} e_{T} &= e_{F} \rho_{0}; \ e_{F} &= e_{T} + e_{C} \rho_{2}; \ e_{C} &= e_{F} \rho_{1} + e_{D} + e_{P}; \\ e_{D} &= e_{C} \rho_{3}; \ e_{P} &= e_{C} \rho_{4} \\ e_{T} &= 1; \ e_{F} &= \frac{1}{\rho_{0}}; \ e_{C} &= \frac{1-\rho_{0}}{\rho_{0}\rho_{2}}; \ e_{D} &= \frac{(1-\rho_{0})\rho_{3}}{\rho_{0}\rho_{2}}; \ e_{P} &= \frac{(1-\rho_{4})\rho_{4}}{\rho_{0}\rho_{2}} \\ \rho_{T} &= \frac{e_{T}}{\lambda} &= \frac{1}{\lambda}; \ \rho_{F} &= \frac{1}{\rho_{0}\mu_{F}}; \ \rho_{C} &= \frac{(1-\rho_{0})}{\rho_{0}\rho_{1}\mu_{C}}; \ \rho_{D} &= \frac{(1-\rho_{0})\rho_{3}}{\rho_{0}\rho_{2}\mu_{0}} \ \rho_{P} &= \frac{(1-\rho_{0})\rho_{4}}{\rho_{0}\rho_{2}\mu_{P}} \\ \pi(n_{T}, n_{F}, n_{C}, n_{D}, n_{p}) &= \frac{1}{G} \frac{(\rho_{T})^{n_{T}}}{n_{T}!} (\rho_{F})^{n_{F}} (\rho_{C})^{n_{C}} (\rho_{D})^{n_{D}} (\rho_{P})^{n_{P}} \end{aligned}$

 $U_F = ?$

• Since U_F is a SSFR, we have

$$U_F = \rho_F \frac{G(5, K-1)}{G(5, K)} = \frac{1}{\rho_0 \mu_F} \frac{G(5, K-1)}{G(5, K)}$$

Average throughput or rate of request completion is:

$$\lambda^* = \mu_F p_0 U_F = rac{G(5, k-1)}{G(5, k)}$$

• but Little's Result ($N = \lambda T$) T = the expected response time = $\frac{K}{\lambda^*} = \frac{KG(5,k)}{G(5,K-1)}$ T = average think + average processing time = $\frac{KG(5,k)}{G(5,K-1)}$ $T = \frac{1}{\lambda}$ + average processing time = $\frac{KG(5,k)}{G(5,K-1)}$ Therefore, average processing time = $\frac{KG(5,k)}{G(5,K-1)} - \frac{1}{\lambda}$

Multiclass Open/Closed/Mixed Jackson Networks

Setting

- We have K customers and M nodes with μ_i exponential service rate for i = 1,.., M.
- *R*, an arbitrary but finite number of classes of customers.
- Let *p_{i,r;j,s}* be the probability that a customer of class *r* completes service at node *i* will become class *s* in node *j*.
- The pairs (*i*, *r*) and (*j*, *s*) belong to the same **subchain** if the **same** customer can visit node *i* in class *r* and node *j* in class *s*.
- Let *m* be the number of subchains, let *E*₁,..., *E_m* be the set of states in each subchains.

Setting: continue

 Let n_{ir} be the number of customers of class r at node i. A "closed" system is characterized by

$$\sum_{i,r)\in E_j} = \text{constant.} \quad \forall j = 1, \dots, m.$$

- For an "*open*" system, a Poisson process with λ_{ir}^0 is the external arrival rate of class *r* to node *i*. Customer may leave the system with $p_{i,r;0}$ so that $\sum_{j,s} p_{i,r;j,s} + p_{i,r;0} = 1$.
- Define $\mathbf{Q}(t) = (\mathbf{Q}_1(t), \dots, \mathbf{Q}_M(t))$ where $\mathbf{Q}_i(t) = (Q_{i1}(t), \dots, Q_{iR}(t))$ with $Q_{ir}(t)$ being the number of class r customers at node i. Note that $\mathbf{Q}(t)$ is **NOT** a CTMC because the class of a customer leaving a node is not known.
- Instead, we define $\boldsymbol{X}_i(t) = (I_{i1}(t), \dots, I_{i,|Q_i(t)|}(t))$ where $I_{ij}(t) \in \{1, 2, \dots, R\}$ is the class of the customer in position *j* in node *i* at time *t*. Then $(\boldsymbol{X}_1(t), \dots, \boldsymbol{X}_M(t)), t \ge 0)$ is a CTMC.

Multiclass Open/Closed/Mixed Jackson Networks: For $k \in \{1, ..., m\}$ such that E_k is an open subchain, let $(\lambda_{ir})_{(i,r) \in E_k}$ be the "unique" strictly positive solution of the traffic equations

$$\lambda_{ir} = \lambda_{ir}^{0} + \sum_{(j,s)\in E_k} \lambda_{js} p_{j,s;i,r} \quad \forall (i,r) \in E_k.$$

For $k \in \{1, ..., m\}$ such that E_k is a closed subchain, let $(\lambda_{ir})_{(i,r)\in E_k}$ be any non-zero solution of

$$\lambda_{ir} = \sum_{(j,s)\in E_k} \lambda_{js} p_{j,s;i,r} \quad \forall (i,r) \in E_k.$$

If $\sum_{r:(i,r)}$ belongs to an open subchain $\lambda_{ir} < \mu_i$ for all i = 1, 2, ..., M, then

$$\pi(\vec{n}) = \frac{1}{G} \prod_{i=1}^{M} \left[n_i! \prod_{r=1}^{R} \frac{1}{n_{ir}!} \left(\frac{\lambda_{ir}}{\mu_i} \right)^{n_{ir}} \right]$$

for all $\vec{n} = ((\vec{n}_1), ..., (\vec{n}_M))$ in the state space, where $(\vec{n}_i) = (n_{i1}, ..., n_{iR}) \in \mathbb{N}^R$ and $n_i = \sum_{r=1}^R n_{ir}$. Here, *G* is the normalization constant.

Example

- There are M = 2 nodes and R = 2 classes of customers. There is no external arrival to node 2. External customers enter node 1 in class 1 with rate λ. Upon completion at node 1, a customer of class 1 is routed to node 2 with the probability 1. Upon completion at node 2, a customer of class a leaves with probability 1.
- There are always *K* customers of class 2 in the system. Upon service completion at node 1 (resp. node 2), customer of class 2 is routed back to node 2 (resp. node 1) in class 2 with probability 1. Let μ_i be the service rate at node *i* = 1,2.

Example: continue

The state space S is

$$S = \{(n_{11}, n_{12}, n_{21}, n_{22}) \in \mathbb{N}^4 : n_{11} \ge 0, n_{21} \ge 0, n_{12} + n_{22} = K\}.$$

There are two subchains: E_1 (open), and E_2 (closed), with $E_1 = \{(1, 1), (2, 1)\}$ and $E_2 = \{(1, 2), (2, 2)\}.$

We find $\lambda_{11} = \lambda_{2,1} = \lambda$ and $\lambda_{12} = \lambda_{22}$. Take $\lambda_{12} = \lambda_{22} = 1$ for instance. We have

$$\pi(\vec{n}) = \frac{1}{G} \left(\frac{n_1}{n_{11}}\right) \left(\frac{n_2}{n_{22}}\right) \left(\frac{\lambda}{\mu_1}\right)^{n_{11}} \left(\frac{\lambda}{\mu_2}\right)^{n_{21}} \left(\frac{1}{\mu_1}\right)^{n_{12}} \left(\frac{1}{\mu_2}\right)^{n_{22}}$$

with $\lambda < \mu_i$, i = 1, 2 (stability condition).

Example: continue (COMPUTING G)

$$\begin{aligned} G &= \sum_{\substack{n_{11} \ge 0; n_{21} \ge 0 \\ n_{12} + n_{22} = K}} \left(\frac{\lambda}{\mu_1}\right)^{n_{11}} \left(\frac{\lambda}{\mu_2}\right)^{n_{21}} \left(\frac{1}{\mu_1}\right)^{n_{12}} \left(\frac{1}{\mu_2}\right)^{n_{22}} \\ &= \left(\sum_{n_{11} \ge 0} \frac{\lambda}{\mu_1}\right)^{n_{11}} \left(\sum_{n_{21} \ge 0} \frac{\lambda}{\mu_2}\right)^{n_{21}} \sum_{n_{12} + n_{22} = K} \left(\frac{1}{\mu_1}\right)^{n_{12}} \left(\frac{1}{\mu_2}\right)^{n_{22}} \\ &= \left(\prod_{i=1}^2 \frac{\mu_i}{\mu_i - \lambda}\right) \left(\frac{1}{\mu_1}\right)^K \sum_{i=1}^K \left(\frac{\mu_1}{\mu_2}\right)^i \\ G &= \left(\prod_{i=1}^2 \frac{\mu_i}{\mu_i - \lambda}\right) \left(\frac{1}{\mu_1}\right)^K \frac{1 - (\mu_1 / \mu_2)^{K+1}}{1 - (\mu_1 / \mu_2)} \quad \text{if } \mu_1 \neq \mu_2, \\ G &= \frac{K + 1}{\mu^K} \left(\frac{\mu}{\mu - \lambda}\right)^2 \qquad \text{if } \mu_1 = \mu_2 = \mu. \end{aligned}$$

Extension to M/M/c

Let $c_i \ge 1$ be the number of servers at node *i* and define $\alpha_i(j) = \min(c_i, j)$ for i = 1, ..., M. Hence $\mu_i \alpha_i(j)$ is the service rate at node *i* when there are *j* customers. We have

$$\pi(\vec{n}) = \frac{1}{G} \prod_{i=1}^{M} \left[\left(\prod_{j=1}^{n_i} \frac{1}{\alpha_i(j)} \right) n_i! \left(\prod_{r=1}^{R} \frac{1}{n_{ir}!} \left(\frac{\lambda_{ir}}{\mu_i} \right)^{n_{ir}} \right) \right].$$

Extension to M/M/c with state-depending routing

Let the total number of customer be $M(\vec{n}) = \sum_{i=1}^{M} n_i$. Let the external arrival rate of class *r* customer at node *i* be $\lambda_{ir}^0 \gamma(M(\vec{n}))$, where γ is an arbitrary function from \mathbb{N} into $[0, \infty)$. We have

$$\pi(\vec{n}) = \frac{d(\vec{n})}{G} \prod_{i=1}^{M} \left[\left(\prod_{j=1}^{n_i} \frac{1}{\alpha_i(j)} \right) n_i! \left(\prod_{r=1}^{R} \frac{1}{n_{ir}!} \left(\frac{\lambda_{ir}}{\mu_i} \right)^{n_{ir}} \right) \right]$$

where

$$d(\vec{n}) = \prod_{j=0}^{M(\vec{n})-1} \gamma(j).$$

and $d(\vec{n}) = 1$ if the network is closed.

A **classic piece** by F. Baskett, K.M. Chandy, R.R. Muntz and F.G. Palacios on "*Open, Closed, and Mixed Networks of Queues with Different Classes of Customers*, JACM, 22(2), 1975.

Terminology

- FCFS: First come first serve M/M/c queue.
- PS: Processor sharing queue.
- LCFS: Last come first serve single server queue.
- IF: Infinite server queue

Characterization

- If node *i* is of type FCFS, then ρ_{ir} = λ_{ir}/μ_i for r = 1,..., R, where R is the number of classes of customer, and μ_i is the exponential service times in node *i*.
- If node *i* is of type PS, LCFS, or IS, then $\rho_{ir} = \lambda_{ir}/\mu_{ir}$ for r = 1, ..., R, and μ_{ir} is the mean service time for customer of type *r* in node *i*.

For nodes of types PS, LCFS, or IS, the service time distribution is **arbitrary**.

 λ_{ir} is the solution of the traffic equations.

Theorem

For a BCMP network with M nodes and R classes of customer, which is open, closed, or mixed in which each node is of type FCFS, PS, LCFS, or IS, the steady state probabilities are:

$$\pi(\vec{n}) = rac{d(\vec{n})}{G} \prod_{i=1}^{M} f_i(\vec{n}_i).$$

where $\vec{n} = (\vec{n}_1, \dots, \vec{n}_M)$ in the state space *S* with $\vec{n}_i = (n_{i1}, n_{i2}, \dots, n_{iR})$ where n_{ir} is the number of jobs of class *r* at node *i*. Moreover, $|\vec{n}_i| = \sum_{r=1}^{R} n_{ir}$ for $i = 1, 2, \dots, M$.

 $G < \infty$ is the normalization constant such that $\sum_{\vec{n} \in S} \pi(\vec{n}) = 1$, $d(\vec{n}) = \prod_{j=0}^{M(\vec{n})-1} \gamma(j)$ if the arrivals depend on the total number of customers $M(\vec{n}) = \sum_{i=1}^{M} |\vec{n}_i|$, and $d(\vec{n}) = 1$ if the network is closed. $f_i(\vec{n}_i)$ for different types of nodes

• If node *i* is of type FCFS:

$$f_i(\vec{n}_i) = |\vec{n}_i|! \prod_{j=1}^{|\vec{n}_i|} \frac{1}{\alpha_i(j)} \prod_{r=1}^R \frac{\rho_{ir}^{n_{ir}}}{n_{ir}!}$$

• If node *i* is of type PS or LCFS:

$$f_i(\vec{n}_i) = |\vec{n}_i|! \prod_{r=1}^R \frac{\rho_{ir}^{n_{ir}}}{n_{ir}!}$$

• If node *i* is of type IS:

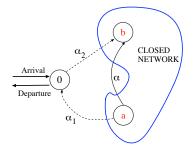
$$f_i(\vec{n}_i) = \prod_{r=1}^R \frac{\rho_{ir}^{n_{ir}}}{n_{ir}!}$$

Comment

- Solve the traffic equations λ_{ir} for i = 1, ..., M and r = 1, ..., R.
- Use queueing network package to solve for G.
- When using queueing network package, all we need to enter are the topology of the network:
 - *K*, the number and the types of nodes,
 - R classes, and
 - routing matrix [p_{i,r;j,s}].
 - external arrival rates
 - service rates, e.g., $\mu_i \alpha_i(j)$, for j = 0, 1, ... for a node that is FCFS.
 - service rates μ_{ir} for customers of class r = 1, ...R for a node that is PS, LCFS, or IS.

Mean value analysis [S. Lavenberg & M. Reiser]

• Derive **expected** performance measures while *avoiding derivation* of steady state probability



• Replace an arc α by $\alpha_1 \rightarrow \text{node } \phi \rightarrow \alpha_2$

- Whenever a customer arrives at node ϕ (via along α_1), it departs from the network and is replaced by a "stochastically" identical customer who has the same routing probability along arc α_2
- This open network behaves exactly as the <u>closed</u> one (except node $\overline{\phi}$)

Mean Value analysis for state-dependent service rates

- The system throughput is still $T(K) = \frac{K}{\sum_{i=1}^{M} v_i W_i(k)}$
- Let π_i(j|K) =Prob[node i has j customers where the network has K customers]

$$W_i(K) = \sum_{j=1}^{K} \pi_i(j-1|K-1) \frac{j}{\mu_i(j)}$$

Example:

$$\pi_i(0|K-1)\frac{1}{\mu_i(1)} + \pi_i(1|K-1)\frac{2}{\mu_i(2)} + \cdots + \pi_i(K-1|K-1)\frac{K}{\mu_i(K)}$$

• The mean queue length is:

$$L_i(K) = \sum_{j=1}^K j\pi_i(j|K)$$

• By definition, $\pi_i(0|0) = 1$ and

$$\pi_i(j|K) = \begin{cases} \frac{\nu_i T(K)}{\mu_i(j)} \pi_i(j-1|K-1) & j=1,2,\ldots,k \\ 1-\sum_{k=1}^K \pi_i(k|K) & j=\phi \end{cases}$$

- For the α we chosen, let us define the network throughput T as the average rate customers pass along arc α in steady state.
- Now we can view *T* as the external arrival rate (γ) in the open network.
- Suppose that we have *M* nodes in the network. Define
 - v_i = average number of visit to node i by a customer = (*visitation rate*)
 - λ_i = average arrival rate of customer to node i

$$\lambda_i = T v_i$$
 , $v_0 = v_a q_{ab}$

• But since customers visit node ϕ EXACTLY ONCE, $v_0 = 1$, therefore:

$$v_a=rac{1}{q_{ab}};$$
 $v_j=\sum_{j=1}^\mu v_j q_{ij}$

• Once one *v_i* is found, we can find other *v_i*'s.

• Looking at node i, let *L_i*= average no. of customers, we have:

$$L_i = \lambda_i w_i$$
, $L_i = T v_i w_i$

• But since $\sum_{i=1}^{M} L_i = K = T \sum_{i=1}^{M} v_i w_i$, therefore the system throughput

$$T = \frac{K}{\sum_{i=1}^{M} v_i w_i}$$

If node i is infinite server, then

$$v_i = \frac{1}{\mu_i}$$

If node i is single server fixed rate (SSFR),

$$w_i = \frac{1}{\mu_i} [Y_i + 1]$$

where Y_i is the mean number of customers seen by an arrival to node i.

John C.S. Lui ()

Example:

$$\pi_i(j|K) = \begin{cases} \frac{\nu_i T(K)}{\mu_i(j)} \pi_i(j-1|K-1) & j=1,2,\ldots,k \\ 1-\sum_{k=1}^K \pi_i(k|K) & j=\phi \end{cases}$$

• Assume it is node 1 (that is , i=1)

$$\pi_{1}(1|1) = \frac{\nu_{1}T(1)}{\mu_{1}(1)}\pi_{1}(\phi|\phi) = \frac{\nu_{1}T(1)}{\mu_{1}(1)}$$

$$\pi_{1}(\phi|1) = 1 - \pi_{1}(1|1)$$

$$\pi_{1}(1|2) = \frac{\nu_{1}T(2)}{\mu_{1}(1)}\pi_{1}(\phi|1)$$

$$\pi_{1}(2|2) = \frac{\nu_{1}T(2)}{\mu_{2}(2)}\pi_{1}(1|1)$$

$$\pi_{1}(\phi|2) = 1 - \pi(1|2) - \pi(2|2)$$